

1 **Transfer entropy and cumulant based cost as**  
2 **measures of nonlinear causal relationships in space**  
3 **plasmas: applications to  $D_{st}$**

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9 **Abstract.** It is well known that the magnetospheric response to the so-  
10 lar wind is nonlinear. Information theoretical tools such as mutual informa-  
11 tion, transfer entropy, and cumulant based analysis are able to characterize  
12 the nonlinearities in the system. Using cumulant based cost, we show that  
13 nonlinear significance of  $D_{st}$  peaks at 3–12 hours lags that can be attributed  
14 to  $VBs$ , which also exhibit similar behavior. However, the nonlinear signif-  
15 icance that peaks at lags 25, 50, and 90 hours can be attributed to internal  
16 dynamics, which may be related to the relaxation of the ring current. These  
17 peaks are absent in the linear and nonlinear self-significance of  $VBs$ . Our  
18 analysis with mutual information and transfer entropy show that both meth-  
19 ods can establish that there are a strong correlation and transfer of infor-  
20 mation from  $V_{sw}$  to  $D_{st}$  at a time scale that is consistent with that obtained  
21 from the cumulant based analysis. However, mutual information also shows  
22 that there is a strong correlation in the backward direction, from  $D_{st}$  to  $V_{sw}$ ,  
23 which is counterintuitive. In contrast, transfer entropy shows that there is  
24 no or little transfer of information from  $D_{st}$  to  $V_{sw}$ , as expected because it  
25 is the solar wind that drives the magnetosphere, not the other way around.  
26 Our case study demonstrates that these information theoretical tools are quite  
27 useful for space physics studies because these tools can uncover nonlinear  
28 dynamics that cannot be seen with the traditional analyses and models that  
29 assume linear relationships.

## 1. Introduction

30 One of the most practically important concepts in dynamical systems is the notion of  
 31 causality. It is particularly useful to organize observational datasets according to causal  
 32 relationships in order to identify variables that drive the dynamics. Understanding causal  
 33 dependencies can also help to simplify descriptions of highly complex physical processes  
 34 because it constrains the coupling functions between the dynamical variables. Analysis  
 35 of those coupling functions can lead to simplification of the underlying physical processes  
 36 that are most important for driving the system. It is particularly useful from a practi-  
 37 cal standpoint to understand causal dependencies in systems involving natural hazards  
 38 because monitoring of causal variables is closely linked with warning.

A common method to establish causal dependencies in a data stream of two variables,  
 e.g.,  $[a(t)]$  and  $[b(t)]$ , is to apply linear correlation studies such as *Strangeway et al.* [2005],  
 which showed the relationship between downward Poynting flux and ion outflows. Causal  
 relationships are typically identified by considering a time-shifted correlation function

$$\lambda_{ab}(\tau) \triangleq \frac{\langle a(t)b(t+\tau) \rangle - \langle a \rangle \langle b \rangle}{\sqrt{\langle a^2 \rangle - \langle a \rangle^2} \sqrt{\langle b^2 \rangle - \langle b \rangle^2}} \quad (1)$$

39 where  $\langle \dots \rangle$  is an ensemble average obtained by drawing samples at a set of measurement  
 40 times,  $\{t_0, t_1, \dots, t_N\}$ . For example, [*Borovsky et al.*, 1998] used such a method to iden-  
 41 tify relationships between solar wind variables and plasma sheet variables. The causal  
 42 dependency that the plasma sheet responds to changes in the solar wind can be identified  
 43 from the time-shift of the peak of the cross correlation indicating a response time. From  
 44 this type of analysis it can be found that the plasma sheet generally responds from the

45 tail to the inner magnetosphere consistent with the notion of earthward convection. Such  
46 analysis has been particularly useful to help understand plasma sheet transport.

47 However, the procedure of detecting causal relationships based on linear cross-  
48 correlation suffers from a number of limitations. First it should be noted that the statisti-  
49 cal accuracy of the correlation function is limited by the resolution and length of the data  
50 stream. Second, the linear time series analysis ignores nonlinear correlations, which may  
51 be important for energy transfer in the magnetospheric system. For example, substorms  
52 are believed to involve storage and release of energy in the magnetotail, which is a highly  
53 nonlinear response. Similarly, magnetosphere-ionosphere coupling may also be highly non-  
54 linear involving the nonlinear development of accelerating potentials along auroral field  
55 lines and nonlinear current-voltage relationships. Third, the cross-correlation may not  
56 be a particularly clear measure when there are multiple peaks or if there is little or no  
57 asymmetry in the forward [i.e.,  $\lambda_{ab}(\tau)$ ] and backward directions [i.e.,  $\lambda_{ba}(\tau) = \lambda_{ab}(-\tau)$ ].  
58 Finally, the cross-correlation does not provide any way to clearly distinguish between two  
59 variables that are passively correlated because of a common driver rather than causally  
60 related.

61 In the remainder of this paper, we will discuss other methods to identify causal rela-  
62 tionships based on entropy based discriminating statistics such as mutual information and  
63 transfer entropy. We will also discuss the cumulant-based method. We will illustrate the  
64 shortcomings and strengths of the various methods for studying causality with examples  
65 from nonlinear dynamics and space physics.

## 2. Linear vs Nonlinear Dependency

66 It is well known that the magnetosphere responds to variation in the solar wind param-  
 67 eters [*Clauer et al.*, 1981; *Baker et al.*, 1983; *Crooker and Gringauz*, 1993; *Papitashvili*  
 68 *et al.*, 2000; *Wing and Johnson*, 2015; *Johnson and Wing*, 2015; *Wing et al.*, 2016], and  
 69 it has been established that the magnetosphere has a significant linear response to the  
 70 solar wind. However, it is also expected that the magnetosphere has a nonlinear response  
 71 [*Tsurutani et al.*, 1990; *Vassiliadis et al.*, 1990; *Klimas et al.*, 1998; *Valdivia et al.*, 2013;  
 72 *Balikhin et al.*, 2011]. The nonlinear response may driven by internal dynamics rather  
 73 than driven externally [*Wing et al.*, 2005; *Johnson and Wing*, 2005]. For example, the  
 74 internal dynamics associated with loading and unloading of magnetic energy associated  
 75 with storms and substorms is nonlinear [e.g., *Johnson and Wing*, 2014, and references  
 76 therein]. Indeed, the data analysis of *Bargatze et al.* [1985] indicated that the dynamical  
 77 response of the magnetosphere to solar wind input could not be entirely understood using  
 78 linear prediction filters.

Suppose that we consider a set of variables  $\mathbf{a}$  and  $\mathbf{b}$  which could be vectors of variables measured in time and we would like to measure their dependency. Instead of considering the covariance matrix/correlation function, we consider a more general measure of dependency between an input and output is obtained by considering whether

$$P(\mathbf{a}, \mathbf{b}) \stackrel{?}{=} P(\mathbf{a})\mathbf{P}(\mathbf{b}). \quad (2)$$

79 where  $P(\mathbf{a}, \mathbf{b})$  is the joint probability of input  $\mathbf{a}$  and output  $\mathbf{b}$  while  $P(\mathbf{a})$  and  $P(\mathbf{b})$  are  
 80 the probability of  $\mathbf{a}$  and  $\mathbf{b}$  respectively. If the relationship holds, then the variables  $\mathbf{a}$   
 81 and  $\mathbf{b}$  are independent. For all other cases, there is some measure of dependency. In the  
 82 case where the system output is completely known given the input,  $P(\mathbf{a}, \mathbf{b}) = \mathbf{P}(\mathbf{a})$ . The

83 advantage of considering Equation 2 is that it is possible to detect the presence of higher  
 84 order nonlinear dependencies between the input and output even in the absence of linear  
 85 dependencies [Gershenfeld, 1998].

## 2.1. Mutual Information and Cumulant based cost

86 Mutual information and cumulant-based cost are two useful measures that quantify  
 87 Eq. 2. Mutual information has the advantage that in the limit of Gaussian joint proba-  
 88 bility distributions, it may be simply related to the correlation coefficient  $C_{ab}(\tau)$  defined  
 89 in equation 1 [Li, 1990]. Cumulants have the advantage of good statistics for limited  
 90 datasets and noisy systems [Deco and Schürmann, 2000]. Moreover, for high-dimensional  
 91 systems it is more efficient to compute moments of the data rather than try to construct  
 92 the probability density function.

Correlation studies also only detect linear correlations, so if the feedback involves non-  
 linear processes (highly likely in this case) then their usefulness may be seriously limited.  
 Alternatively, entropy-based measures such as mutual information [Prichard and Theiler,  
 1995; Materassi et al., 2011] and cumulants [Johnson and Wing, 2005] are useful for de-  
 tecting linear as well as nonlinear correlations. The mutual information is constructed  
 from the probability distribution function of the variables and may be computed using  
 an quantization procedure where data is binned such that the samples  $[a(t)]$  are assigned  
 discrete values  $\hat{a} \in \{a_1, a_2, \dots, a_n\}$  of an alphabet  $\aleph_1$  and  $[b(t)]$  is assigned discrete values  
 $\hat{b} \in \{b_1, b_2, \dots, b_m\}$  of an alphabet  $\aleph_2$ . The *ad hoc* time-shifted mutual entropy

$$\mathcal{M}_{ab}(\tau) \triangleq \sum_{\hat{a} \in \aleph_1, \hat{b} \in \aleph_2} p(\hat{a}(t + \tau), \hat{b}(t)) \log \left( \frac{p(\hat{a}(t + \tau), \hat{b}(t))}{p(\hat{a})p(\hat{b})} \right) \quad (3)$$

93 has been used as an indicator of causality, but suffers from the same problems as time-  
 94 shifted cross correlation when it has multiple peaks and long range correlations.

Similarly, examination of time-shifted cumulants could be used as an indicator of causal-  
 ity in a nonlinear system. In this case, we can define a discriminating statistic

$$D^C = \sum_{q=1}^{\infty} \sum_{i_1, \dots, i_q \in \Pi_q} K_{1i_2 \dots i_q}^2 \quad (4)$$

where

$$\begin{aligned} K_i &= C_i = \langle z_i \rangle & (4) \\ K_{ij} &= C_{ij} - C_i C_j = \langle z_i z_j \rangle - \langle z_i \rangle \langle z_j \rangle \\ K_{ijk} &= C_{ijk} - C_{ij} C_k - C_{jk} C_i - C_{ik} C_j + 2C_i C_j C_k \\ K_{ijkl} &= C_{ijkl} - C_{ijk} C_l - C_{ijl} C_k - C_{ilk} C_j - C_{ljk} C_i \\ &\quad - C_{ij} C_{kl} - C_{il} C_{kj} - C_{ik} C_{jl} + 2(C_{ij} C_k C_l \\ &\quad + C_{ik} C_j C_l + C_{il} C_j C_k + C_{jk} C_i C_l + C_{jl} C_i C_k \\ &\quad + C_{kl} C_i C_j) - 6C_i C_j C_k C_l \end{aligned}$$

are the cumulants

$$C_{i \dots j} = \int d\mathbf{z} P(\mathbf{z}) z_i \dots z_j \equiv \langle z_i \dots z_j \rangle \quad (5)$$

95 of the joint probability distribution for variables  $z_1, \dots, z_j$ .

With only two variables,  $a$  and  $b$ , defined above, we can consider the cost function

$$D_{a,b}^C(\tau) = D_C(a(t), b(t + \tau)) \quad (6)$$

96 The presence of nonlinear dependence has been identified by comparing the cumulant cost  
 97 for a time series with the cumulant based cost of surrogate time series, which are con-  
 98 structed to have the same linear correlations as in [*Johnson and Wing, 2005*]). Significance  
 99 measures the difference in the discriminating statistic from the mean of the discriminating  
 100 statistic of the surrogates in terms of the spread of the surrogates,  $\sigma$ .

101 In Section 3, we will show an application of cumulant based analysis to the distur-  
 102 bance storm-time index ( $D_{st}$ ). In principle, the cross-correlation, mutual information,  
 103 and cumulant-based cost should be independent of the selection of measurement points

104 if the system is stationary; therefore, time stationarity can be examined by comparing  
 105 these discriminating statistics for groups of measurements drawn from different windows  
 106 of time as in [Johnson and Wing, 2005; Wing et al., 2016].

## 2.2. Transfer entropy

Another method for determining causality is the one-sided transfer entropy [Schreiber, 2000; De Michelis et al., 2011; Materassi et al., 2014; Wing et al., 2016, 2018], which is based upon the conditional mutual information

$$\mathcal{M}_C(x, y|z) \triangleq \sum_{x \in \mathbb{N}_1} \sum_{y \in \mathbb{N}_2} \sum_{z \in \mathbb{N}_3} p(x, y, z) \log \left( \frac{p(x, y, z)p(z)}{p(x, z)p(y, z)} \right) \quad (7)$$

107 The conditional mutual information measures the dependence of two variables,  $x$  and  $y$ ,  
 108 given a conditioner variable,  $z$ . If either  $x$  or  $y$  are dependent on  $z$  the mutual information  
 109 between  $x$  and  $y$  is reduced, and this reduction of information provides a method to  
 110 eliminate coincidental dependence, or conversely to identify causal dependence.

Transfer entropy considers the conditional mutual information between two variables using the past history of one of the variables as the conditioner.

$$\mathcal{T}_{a \rightarrow b}(\tau) = \sum_{\hat{a} \in \mathbb{N}_1} \sum_{\hat{a}^{(k)} \in \mathbb{N}_1^{(k)}} \sum_{\hat{b} \in \mathbb{N}_2} p(\hat{a}(t + \tau), \hat{a}^{(k)}(t), \hat{b}(t)) \log \left( \frac{p(\hat{a}(t + \tau)|\hat{a}^{(k)}(t), \hat{b}(t))}{p(\hat{a}(t + \tau)|\hat{a}^{(k)}(t))} \right) \quad (8)$$

111 where  $\hat{a}^{(k)}(t) = [\hat{a}(t), \hat{a}(t - \Delta), \dots, \hat{a}(t - (k - 1)\Delta)]$ . The standard definition of transfer  
 112 entropy takes  $k = 1$  (no lag), but keeping a higher embedding dimension could in prin-  
 113 ciple provide a more precise measure (for example, if  $a$  has periodicity a dimension of 2  
 114 may provide better prediction of future values of  $a$  from its past time series and therefore  
 115 lower the transfer entropy. Transfer entropy as a discriminating statistic has the following  
 116 advantages. First in the absence of information flow from  $a$  to  $b$  (i.e.,  $a(t + \tau)$  has no  
 117 additional dependence from  $b(t)$  beyond what is known from the past history of  $a^{(k)}(t)$ )



118  $p(\hat{a}(t + \tau) | \hat{a}^{(k)}(t), \hat{b}(t)) = p(\hat{a}(t + \tau) | \hat{a}^{(k)}(t))$  and the transfer entropy vanishes. The transfer  
 119 entropy is also highly directional so that  $\mathcal{T}_{a \rightarrow b} \neq \mathcal{T}_{b \rightarrow a}$ . The advantage can be clearly  
 120 seen for dynamical systems where variables are forward differenced and the transfer en-  
 121 tropy is clearly one-sided while mutual information and correlation functions can even be  
 122 symmetric [Schreiber, 2000]. This measure also accounts for static internal correlations,  
 123 which can be used to determine whether two variables are driven by a common driver or  
 124 whether the variable  $b$  is causally driving the variable  $a$ .

125 Both mutual information and transfer entropy require binning of data. As mentioned  
 126 in Wing et al. [2016], the number of bins ( $n_b$ ) needs to be chosen properly and there are  
 127 some guidelines that can be followed. In general, we would like to maximize the amount  
 128 of information. Having too few bins would lump too many points into the same bin,  
 129 leading to loss of information. Conversely, having too many bins would leave many bins  
 130 with 0 or a few number of points, which also would lead to loss of information. Sturges  
 131 [1926] proposed that for a normal distribution, optimal  $n_b = \log_2(n) + 1$  and bin width  
 132 ( $w$ ) =  $range/n_b$ , where  $n$  = number of points in the dataset,  $range$  = maximum value -  
 133 minimum value of the points. In practice, there is usually a range of  $n_b$  that would work.

### 3. Application to space weather: $D_{st}$ analysis

134  $D_{st}$  (disturbance storm time index) is an hourly index that gives a measure of the  
 135 strength of the symmetric ring current that, in turn, provides a measure of the dynamics  
 136 of geomagnetic storms [Dessler and Parker, 1959]. Because of its global nature,  $D_{st}$  is  
 137 often used as one of the several indices that represent the state of the magnetosphere.  
 138 For example, Balasis et al. [2011] used the cumulative square amplitude of  $D_{st}$  time series  
 139 as a proxy for energy dissipation rate in the magnetosphere and found that it fits well

140 a power law with log-periodic oscillations, which was interpreted as evidence for discrete  
141 scale invariance in the  $D_{st}$  dynamics.

142 When plasma sheet ions are injected into the Earth inner magnetosphere, they drift  
143 westward around the Earth, forming the ring current. Studies have shown that the  
144 substorm occurrence rate increases with solar wind velocity (high speed streams) [e.g.,  
145 *Kissinger et al.*, 2011; *Newell et al.*, 2016]. An increase in the solar wind electric field,  
146  $VB_z$ , can increase the dawn-dusk electric field in the magnetotail, which in turn deter-  
147 mines the amount of plasma sheet particles that move to the inner magnetosphere [e.g.,  
148 *Friedel et al.*, 2001]. Studies have shown that the electric field,  $VB_s$  ( $V_{sw} \times$  southward  
149 IMF  $B_z$ ) or  $VB_z$ , has a strong effect on the ring current dynamics [*Burton et al.*, 1975;  
150 *O'Brien and McPherron*, 2000; *McPherron and O'Brien*, 2001; *Weygand and McPherron*,  
151 2006].

152 For the present study, we examine the relationships between solar wind velocity ( $V_{sw}$ )  
153 and  $VB_s$  with  $D_{st}$ . We use  $D_{st}$  records in the period 1974 – 2001 obtained from  
154 Kyoto University World Data Center for Geomagnetism ([http://swdcwww.kugi.kyoto-](http://swdcwww.kugi.kyoto-u.ac.jp/index.html)  
155 [u.ac.jp/index.html](http://swdcwww.kugi.kyoto-u.ac.jp/index.html)). The corresponding solar wind data are obtained from IMP-8, ACE,  
156 WIND, ISEE1, and ISEE3 observations. The ACE SWEPAM and MAG data; and  
157 the WIND MAG data are obtained from CDAWeb (<http://cdaweb.gsfc.nasa.gov/>). The  
158 WIND 3DP data are obtained from the 3DP team directly. The ISEE1 and ISEE3  
159 data are obtained from UCLA (these datasets are also available at NASA NSSDC  
160 [<http://nssdc.gsfc.nasa.gov/space/>]). The IMP8 data come directly from the IMP teams.  
161 The solar wind is propagated with minimum variance technique [*Weimer et al.*, 2003] to

162 GSM  $(X, Y, Z) = (17, 0, 0) R_E$  to produce 1-min files, from which hourly averaged solar  
 163 wind parameters are constructed.

### 3.1. Cumulant based analysis

Section 2.1 presents the method of cumulant based cost. Here, we show an application of cumulant based cost to detect nonlinear dynamics in  $D_{st}$ . We consider the forward coupling between a solar wind variable such as  $VB_s$  and  $D_{st}$ , which characterizes the ring current response to the solar wind driver. We therefore consider the nonlinear cross-correlations of the vector

$$\mathbf{c}(t, \tau) = \{VB_s(t), D_{st}(t + \tau)\} = \{z_1, z_2\} \quad (9)$$

164 The generalization of cost is based on realizations of  $\{z_1, z_2\}$ . In this case, each variable  
 165 is Gaussianized with unit variance to eliminate static nonlinearities (i.e. higher order  
 166 self-correlations in  $VB_s$  and  $D_{st}$  are eliminated so that the cost measures only cross-  
 167 dependence between  $VB_s$  and  $D_{st}$ ). This procedure is explained in the next paragraph.

168 The distribution of  $D_{st}$  and  $VB_s$  are generally non-Gaussian. As such, the raw dis-  
 169 tributions (e.g., distribution of values of  $D_{st}$ ) may have nonzero higher-order cumulants  
 170 (e.g., they can have a skew and kurtosis). This property makes it more difficult to in-  
 171 terpret whether the higher order cumulants in the time evolution arise from the overall  
 172 shape of the distribution of data points or from the time-ordering of the data. To elim-  
 173 inate the inherent nonzero cumulants in the overall distribution of data, we construct a  
 174 rank-ordered map from the original dataset to a proxy dataset of the same length drawn  
 175 from a Gaussian distribution [Kennel and Isabelle, 1992; Schreiber and Schmitz, 1996;  
 176 Deco and Schürmann, 2000]. The distribution of the proxy dataset ensures that all cu-

177 cumulants of the distribution beyond second order should in principle vanish. However, the  
 178 time-ordering of the data can still lead to nonzero cumulants, because the joint probability  
 179 distribution of  $D_{st}(t+\tau)$  and  $D_{st}(t)$  may be non-Gaussian even if the distribution of  $D_{st}$  is  
 180 Gaussian. Moreover, it is simple to construct surrogate data from the Gaussianized data  
 181 that shares the same autocorrelation by using the same power spectrum, but randomly  
 182 shifting the phases of the Fourier coefficients. The surrogate data therefore has the same  
 183 autocorrelation as the original data. Any deviation from the linear statistic is apparent  
 184 from comparison with the surrogate data, and we interpret these deviations as evidence  
 185 of nonlinear dependence because we have falsified the hypothesis that the data can be  
 186 adequately described by linear statistics. This method has been successfully employed in  
 187 *Johnson and Wing* [2005] where  $K_p$  record was analyzed with mutual information and  
 188 cumulants.

189 In Figure 1 we plot the significance obtained from the year 1999 as a function of time  
 190 delay,  $\tau$ . Significance extracted from  $\{VBs(t), D_{st}(t + \tau)\}$  and  $\{VBs(t), VBs(t + \tau)\}$   
 191 for 1999 are plotted in panels (a) and (b), respectively. It should be noted that there  
 192 is a strong linear response at around 3 hour time delay. As shown in Figure 1a, there  
 193 is a clear nonlinear response with peaking around 3–10, 25, 50 and 90 hours lasting for  
 194 approximately 1 week. In contrast, in Figure 1b, the nonlinearity only has one broad peak  
 195 around 3 – 12 hours in the self-significance for  $VBs$ , suggesting that the nonlinear and  
 196 linear peaks at  $\tau = 3$ –12 hours in in Figure 1a i may be associated with  $VBs$ . We will  
 197 revisit the solar wind causal relationship with  $D_{st}$  using transfer entropy in Section 3.2.

198 The absence of the nonlinear peaks at  $\tau = 25, 50,$  and  $90$  hours in the self-significance  
 199 for  $VBs$  (Figure 1b) suggest that these nonlinearities in  $\{VBs(t), D_{st}(t+\tau)\}$  are related to

200 internal magnetospheric dynamics. As the  $D_{st}$  index is thought to reflect storm activity,  
201 it is reasonable that nonlinear significance would decay on the order of 1 week as storms  
202 commonly last around that time. The strong nonlinear responses at  $\tau = 25, 50,$  and 90  
203 hours are likely related to multiple modes of relaxation of the ring current following the  
204 commencement of storms. It should also be noted that other nonlinearities detected by  
205 even higher order cumulants may also be present; however, the calculation demonstrates  
206 the nonlinear nature of the underlying dynamics.

207 A common scenario for storm-ring current interaction is the following. A storm com-  
208 presses the magnetosphere and intensifies the magnetic field in the magnetosphere and  
209 energetic particles are injected into the ring current region. The ring current intensifies  
210 as more particles are injected during the main phase of the storm, which can last  $\sim 6$   
211 hours [Weygand and McPherron, 2006]. Once the injection stops, the ring current begins  
212 to decay and the storm enters the recovery phase. Conservation of magnetic moment  
213 implies that anisotropies develop in the ring current and plasma sheet. Anisotropy drives  
214 the ring current plasma unstable to ion cyclotron waves. The ion cyclotron waves scatter  
215 energetic ions into the loss cone so that they are lost from the ring current. Nonlinear  
216 interaction between waves and particles keeps the plasma near marginal instability with a  
217 steady loss of energetic particles due to wave-particle scattering. Other loss mechanisms  
218 include charge exchange, coulomb scattering, and convective of ions to the front of the  
219 magnetopause. The ring current decay can have two stages [Kozyra et al., 2002]. In the  
220 first stage, the ring current decays rapidly and the loss mechanisms can be attributed to  
221 convective out flow, pitch-angle scattering in the ring current, and  $O^+$  charge exchange  
222 [e.g., Weygand and McPherron, 2006; Hamilton et al., 1988]. The second stage may typi-

223 cally begin about one day from the commencement of the storm (see, for example, Figure  
 224 7 of *Kozyra et al.* [2002]). In the second stage, the decay rate is slower and is attributed  
 225 mainly to  $H^+$  charge exchange [*Hamilton et al.*, 1988] and can take several days to de-  
 226 plete the ring current to the baseline level [*Smith et al.*, 1976]. We can speculate that  
 227 the multiple nonlinear response lag times that are detected with the cumulant-based ap-  
 228 proach are likely the relaxation of the ring current due to complex interplay of multiple  
 229 loss processes.

### 3.2. Transfer entropy

230 As mentioned in Section 2.2, transfer entropy gives a measure of how much information  
 231 is transferred from one variable to another. We have applied transfer entropy and mutual  
 232 information to the relationship between the  $V_{sw}$  and  $D_{st}$  for the period 1974 – 2001. The  
 233 result is shown in Figure 2. Note that the mutual information measure suggests strong  
 234 correlations between prior values of  $D_{st}$  and  $V_{sw}$ . This finding suggests that  $D_{st}$  could be  
 235 a driver of  $V_{sw}$ , which is counterintuitive. On the other hand, the transfer entropy clearly  
 236 shows that this information transfer in the backward direction ( $D_{st} \rightarrow V_{sw}$ ) does not rise  
 237 above the noise level (the horizontal blue lines indicate mean and standard deviation of  
 238 100 surrogate data sets where the data was randomly reordered.) This result is expected  
 239 because it is the solar wind that drives the magnetosphere, not the other way around.  
 240 The transfer of information from  $V_{sw}$  to  $D_{st}$  peaks at  $\tau = 8 - 11$  hours. The cumulant  
 241 based analysis in Section 3.1 shows that the response of  $D_{st}$  to  $VBs$  has similar time scale.  
 242 This time scale is consistent with the 4 to 15 hours transport time for the solar wind to  
 243 reach the midnight and noon regions of the geosynchronous orbit, respectively, from the

244 dayside magnetopause [Borovsky et al., 1998]. The analysis presented here illustrates the  
245 power of the transfer entropy for accessing causality.

#### 4. Summary

246 We recently used mutual information, transfer entropy, and conditional mutual infor-  
247 mation to discover the solar wind drivers of the outer radiation belt electrons [Wing et al.,  
248 2016]. Because  $V_{sw}$  anticorrelates with solar wind density ( $n_{sw}$ ), it is hard to isolate the  
249 effects of  $V_{sw}$  on radiation belt electrons, given  $n_{sw}$  and vice versa. However, using condi-  
250 tional mutual information, we were able to determine the information transfer from  $n_{sw}$   
251 or any other solar wind parameters to radiation belt electrons, given  $V_{sw}$  (or any other  
252 solar wind parameters). We also showed that the triangle distribution in the radiation  
253 belt electron vs. solar wind velocity plot [Reeves et al., 2011] can be understood better  
254 when we consider that  $V_{sw}$  and  $n_{sw}$  transfer information to radiation belt electrons with  
255 2 days and 0 day ( $< 24$  hr) lags, respectively. Also recently, we used transfer entropy to  
256 better understand the causal parameters in the solar cycle and their response lag times  
257 [Wing et al., 2018].

258 As a follow up to Wing et al. [2016, 2018], the present study demonstrates further how  
259 information theoretical tools can be useful for space physics and space weather studies.  
260 Cumulant based analysis can be used to distinguish internal vs. external driving of the  
261 system. Both mutual information and transfer entropy give a measure of shared infor-  
262 mation between two variables (or vectors). However, unlike mutual information, transfer  
263 entropy is highly directional. To illustrate, we apply mutual information, transfer entropy,  
264 and cumulant based analysis to investigate the dynamics of  $D_{st}$  index.

Our analysis with mutual information and transfer entropy indicates that there are strong linear and nonlinear correlations and transfer of information, respectively, in the forward direction between  $V_{sw}$  and  $D_{st}$  ( $V_{sw} \rightarrow D_{st}$ ). However, mutual information indicates that there is also a strong correlation in the backward direction ( $D_{st} \rightarrow V_{sw}$ ), which is puzzling and counterintuitive. In contrast, the transfer entropy indicates that there is no information transfer in the backward direction ( $D_{st} \rightarrow V_{sw}$ ), as expected because it is the solar wind that drives the magnetosphere, not the other way around. The transfer of information from  $V_{sw}$  to  $D_{st}$  peaks at  $\tau = 8 - 11$  hours.

Using the cumulant-based significance, we have established that the underlying dynamics of  $D_{st}$  is in general nonlinear exhibiting a quasiperiodicity which is detectable only if nonlinear correlations are taken into account. The strong nonlinear responses of  $D_{st}$  to  $VBs$  at  $\tau = 25, 50,$  and  $90$  hours are likely related to multiple modes of relaxation of the ring current from multiple loss mechanisms following the commencement of storms. It is, of course, possible that these nonlinearities are caused by solar wind drivers other than  $VBs$ . However, the timing of these nonlinearities would put them well in the recovery phase of a storm and previous studies suggested that the ring current decays in the recovery phase are strongly influenced by  $VBs$  [Burton et al., 1975; O'Brien and McPherron, 2000; McPherron and O'Brien, 2001]. The nonlinearities at  $\tau = 3 - 12$  hours are not caused by internal dynamics but rather by the solar wind driver, which is similar with the time scale for the solar wind transport time from the dayside magnetopause to the inner magnetosphere. This time scale is consistent with the time scale for the information transfer from the solar wind to  $D_{st}$  obtained from transfer entropy analysis.



287 Although linear models are useful, our results indicate that these models have to be  
288 used with cautions because solar wind – magnetosphere system is inherently nonlinear.  
289 Hence, nonlinearities generally need to be taken into account in order to describe the  
290 system accurately. Local-linear models (which include slow evolution of parameters) may  
291 be able to handle some nonlinearities, but it is expected that these local-linear models  
292 would have difficulties if the dynamics suddenly and rapidly change.

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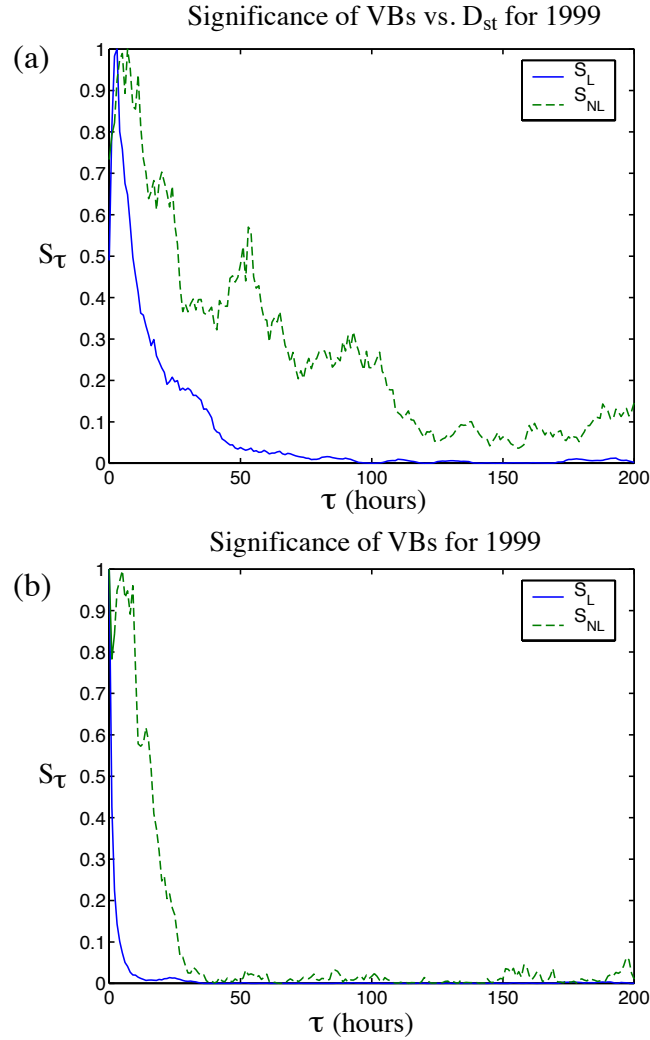
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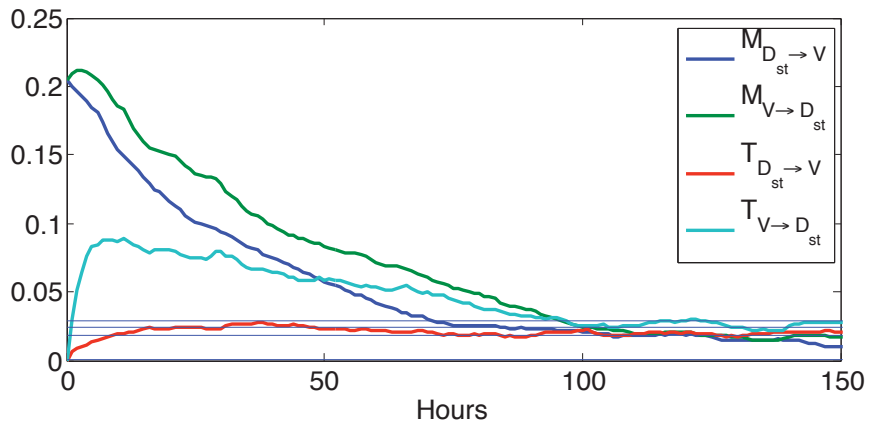
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**Figure 1.** Significance extracted from (a)  $\{VBs(t), D_{st}(t-\tau)\}$  and (b)  $\{VBs(t), VBs(t-\tau)\}$  for 1999. It should be noted that there is a strong linear response at around 3 hour time delay. There is a clear nonlinear response with a strong peak around 50 hours lasting for approximately 1 week. The longterm nonlinear response is absent in the solar wind data indicating that the longterm nonlinear correlations between  $VBs$  and  $D_{st}$  are the result of internal magnetospheric dynamics.





**Figure 2.** Comparison of mutual information and transfer entropy measures to determine causal driving of the magnetosphere as characterized by  $D_{st}$ . Note that causal driving appears to peak somewhat later (11 hours) than indicated by mutual information (2 hours) indicating that internal dynamics likely are very important initially. The backward transfer entropy is below the noise level for all values indicating that  $D_{st}$  in no way influences the upstream solar wind velocity. Such a conclusion could not be inferred from the mutual information measure.